# General Certificate of Education (A-level) June 2012 

## Statistics

SS05

## (Specification 6380)

## Statistics 5

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## Key to mark scheme abbreviations

| M | mark is for method |
| :--- | :--- |
| m or dM | mark is dependent on one or more M marks and is for method |
| A | mark is dependent on M or m marks and is for accuracy |
| B | mark is independent of M or m marks and is for method and accuracy |
| E | mark is for explanation |
| Jor ft or F | follow through from previous incorrect result |
| CAO | correct answer only |
| CSO | correct solution only |
| AWFW | anything which falls within |
| AWRT | anything which rounds to |
| ACF | any correct form |
| AG | answer given |
| SC | special case |
| OE | or equivalent |
| A2,1 | 2 or 1 (or 0) accuracy marks |
| $-x$ EE | deduct $x$ marks for each error |
| NMS | no method shown |
| PI | possibly implied <br> SCA |
| substantially correct approach |  |
| cf | candidate |
| dp | significant figure(s) |
| decimal place(s) |  |

## No Method Shown

Where the question specifically requires a particular method to be used, we must usually see evidence of use of this method for any marks to be awarded.

Where the answer can be reasonably obtained without showing working and it is very unlikely that the correct answer can be obtained by using an incorrect method, we must award full marks. However, the obvious penalty to candidates showing no working is that incorrect answers, however close, earn no marks.

Where a question asks the candidate to state or write down a result, no method need be shown for full marks.
Where the permitted calculator has functions which reasonably allow the solution of the question directly, the correct answer without working earns full marks, unless it is given to less than the degree of accuracy accepted in the mark scheme, when it gains no marks.

Otherwise we require evidence of a correct method for any marks to be awarded.

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 1(a) | $\mathrm{s}^{2}=860.4$ | B1 | 1 | B1 860.4 (860~861) |
| (b) | $90 \%$ confidence interval of s.d. given by |  |  |  |
|  | $3.325<9 \times 860.4 / \sigma^{2}<16.919$ | m1 |  | generous, <br> allow slip $\left(10 \times \mathrm{s}^{2} \quad 9 \times \mathrm{s}\right)$ <br> m 1 completely correct expression <br> - allow incorrect $\chi^{2}$ values |
|  | 7743.6/16.919 < $\sigma^{2}<7743.6 / 3.325$ | B1 |  | B19df |
|  |  | B1 |  | B1 3.325 and 16.919 |
|  | $457.687<\sigma^{2}<2328.902$ | M1 |  | M1 correct method for interval for $\sigma$ (or $\sigma^{2}$ provided it is clearly called $\sigma^{2}$ or variance) |
|  | $21.4<\sigma<48.3$ | A1 | 6 | $\begin{aligned} & \text { A1 } 21.4 \text { (21.35~21.45) and } \\ & \quad 48.3 \text { (48.2~48.3) } \end{aligned}$ |
| (c) | 60 mm is above the upper limit of | E1 |  | E1 above confidence interval |
|  | be necessary to allow for such a large standard deviation. | E1」 | 2 | E1 $\checkmark$ unnecessary |
|  | Total |  | 9 |  |


| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 2(a)(i) | mean 15 | B1 |  | B1 15 cao |
|  | s.d. $30 / \sqrt{ } 12=8.66$ | $\begin{gathered} \text { M1 } \\ \text { A1 } \end{gathered}$ | 3 | M1 method for s.d. or variance A1 8.66 (8.65~8.7) |
| (ii) | $18 / 30=0.6$ | $\begin{gathered} \text { M1 } \\ \text { A1 } \end{gathered}$ | 2 | M1 method - allow wrong tail A1 0.6 acf |
| (b)(i) | $\mathrm{z}=(12-10) / 3.1=0.645$ | M1 |  | M1 method - allow wrong tail |
|  | $\mathrm{P}(>12)=1-0.741=0.259$ | A1 | 2 | A1 0.259 (0.257~0.262) |
| (ii) | Alan's waiting time is shorter on average and also less variable. His probability of having to wait more than 12 minutes is much less than Megara's | E1 E1 | 2 | E1 average wait shorter <br> E1 less variable <br> E1 prob $>12$ much less maximum 2 |
| (c) | Megara's waiting time is now rectangular on $[0,20]$ mean 10 s.d. $20 / \sqrt{ } 12=5.77$ | M1 A1 | 2 | M1 rectangular [0,20] may be implied A1 10 and 5.77 (5.75~5.8) |
|  | Total |  | 11 |  |
| 3(a)(i) | $\begin{aligned} \text { mean } & =1 / 0.0045 \\ & =222.2 \end{aligned}$ | $\begin{gathered} \hline \text { M1 } \\ \text { A1 } \end{gathered}$ | 2 | $\begin{aligned} & \hline \text { M1 method } \\ & \text { A1 } 222(222 \sim 222.4) \end{aligned}$ |
| (ii) | probability will wear the suit in next 100 days $\begin{aligned} & =1-\mathrm{e}^{-0.45} \\ & =1-0.638=0.362 \end{aligned}$ | $\begin{aligned} & \text { M1 } \\ & \text { m1 } \\ & \text { A1 } \end{aligned}$ | 3 | M1 $100 \times 0.0045$ m 1 method - allow wrong tail A1 0.362 ( $0.362 \sim 0.363$ ) |
| (iii) | probability will not wear suit for a $\begin{aligned} \text { year } & =\mathrm{e}^{-365 \times 0.0045} \\ & =\mathrm{e}^{-1.6425} \\ & =0.193 \end{aligned}$ | M1 A1 | 2 | M1 method - allow wrong tail A1 0.193 (0.193~0.194) |
| (b) | $\begin{aligned} \text { mean } & =365 \times 0.0045 \\ & =1.64 \end{aligned}$ | $\begin{gathered} \text { M1 } \\ \text { A1 } \end{gathered}$ | 2 | $\begin{aligned} & \text { M1 method } \\ & \text { A1 } 1.64(1.64 \sim 1.65) \end{aligned}$ |
| (c) | number of times per year which Imran wears a suit is Poisson mean $1.64+1.72=3.36$ | $\begin{aligned} & \text { B1 } \\ & \text { B1 } \end{aligned}$ | 2 | $\begin{aligned} & \text { B1 Poisson, mean } 1.72+\text { their (b) } \\ & \text { B1 } 3.36(3.36 \sim 3.37) \end{aligned}$ |
|  | Total |  | 11 |  |


| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 4(a)(i) | $\begin{aligned} & \mathrm{s}_{1}{ }^{2}=3742.49\left(\mathrm{~s}_{1}=61.18\right) \\ & \mathrm{s}_{2}{ }^{2}=4716.14\left(\mathrm{~s}_{2}=68.67\right) \end{aligned}$ | B1 |  | B1 $3742.49(3740 \sim 3745)$ and $4716.14(4710 \sim 4720)$ |
|  | $\begin{aligned} & \mathrm{H}_{0}: \sigma_{1}=\sigma_{2} \\ & \mathrm{H}_{1}: \sigma_{1} \neq \sigma_{2} \end{aligned}$ | B1 |  | B1 hypotheses correct |
|  | $\mathrm{F}=4716.14 / 3742.49=1.26$ | $\begin{aligned} & \text { M1 } \\ & \text { A1 } \end{aligned}$ |  | M1 method for F <br> A1 1.26 ( $1.255 \sim 1.265$ ) <br> or 0.794 (0.793-0.794) |
|  | c.v. $\mathrm{F}_{[6,9]}$ is 4.32 | $\begin{gathered} \mathrm{B} 1 \\ \text { B1 } \downarrow \end{gathered}$ |  | $\begin{aligned} & \text { B1 6,9 df } \\ & \text { B1 } \sqrt{ } 4.32 \text { - their df } \\ & \text { [Or 0.794 (0.793-0.794); 9,6df; } \\ & 0.231] \end{aligned}$ |
|  | Accept $\mathrm{H}_{0}$, no significant evidence that standard deviation has changed after October 2011 | A1 | 7 | A1 accept $\mathrm{H}_{0}$ must be compared with F <br> (or $\mathrm{p}=0.7245$ compared with 0.05 ) |
| (ii) | $\bar{X}_{1}=648.6 \quad \bar{X}_{2}=619.86$ | B1 |  | B1 $648.6(648 \sim 649)$ and 619.86 <br> (619.5 ~ 620) |
|  | Pooled variance estimate $\begin{aligned} & \mathrm{s}_{\mathrm{p}}^{2}=(3742.49 \mathrm{x} 9+4716.14 \mathrm{x} 6) / 15 \\ & =4131.95 \end{aligned}$ | M1 |  | M1 method for pooled variance |
|  | $\mathrm{H}_{0}: \mu_{1}=\mu_{2}$ | B1 |  | B1 one hypothesis correct |
|  | $\mathrm{H}_{1}: \mu_{1}>\mu_{2}$ | B1 |  | B1 both hypotheses correct - don't penalise the same error twice |
|  | $t=\frac{(648.6-619.86)}{\sqrt{4131-95(1 / 10+1 / 7)}}$ | M1 <br> M1 |  | M1 method for numerator M1 method for denominator - |
|  | $=0.907$ | A1 |  | A1 0.907 ( $0.9 \sim 0.91$ ) - ignore sign |
|  | c.v. $\mathrm{t}_{15}$ is 1.753 | B1 |  | B1 15 df . |
|  |  | B1 |  | B1 1.753 - ignore sign |
|  | Accept $\mathrm{H}_{0}$ i.e. no significant evidence of a reduction in Saturday takings after October 2011 | $\begin{aligned} & \mathrm{A} 1 \checkmark \\ & \mathrm{~A} 1 \checkmark \end{aligned}$ | 11 | A1 $\sqrt{\text { accept }} \mathrm{H}_{0}$ - must be compared with correct tail of $t$ Al $\sqrt{ }$ conclusion in context 0 for contradiction |
|  |  |  |  | ( or $\mathrm{p}=0.189$ compared with 0.05 ) |


| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 4(b)(i) | $\begin{aligned} & \mathrm{H}_{0}: \mu_{2}=\mu_{1}+50 \\ & \mathrm{H}_{1}: \mu_{2}>\mu_{1}+50 \end{aligned}$ | $\begin{aligned} & \text { B1 } \\ & \text { B1 } \end{aligned}$ | 2 | B1 1 correct hypothesis B1 both correct - only penalise the same mistake once |
| (ii) | 801,887,1013,884,964,1014,1146 | $\begin{gathered} \text { M1 } \\ \text { A1 } \end{gathered}$ | 2 | M1 method <br> A1 accuracy - allow one slip |
| (b)(iii) | critical value of $\mathrm{t}_{15}$ is 2.602 | B1 |  | B1 2.602 |
|  | reject $\mathrm{H}_{0}$, conclude total takings will be increased by more than $£ 50$. | $\begin{aligned} & \text { B1 } \\ & \text { B1 } \end{aligned}$ | 3 | B1 conclusion (M implied) B1 in context must be compared with $t$ - values $\text { (or } \mathrm{p}=0.0000936 \text { ) }$ |
| (c) | There is no significant evidence that Saturday takings have been reduced and there is significant evidence that | E1 |  | E1 Saturday takings not reduced |
|  | total weekend takings have increased by more than $£ 50$ per week. However the conclusions should be treated | E1 |  | E1 no change in variability of Saturday takings |
|  | with caution because the samples of weekends are not random and in particular the takings after Sunday | E1 |  | E1 Total weekend takings increased more than $£ 50$ (maximum 2) |
|  | approach of Christmas. Sunday takings are increasing steadily perhaps due to Christmas or customers getting used to Sunday opening. | E1 | 4 | E1 samples not random <br> E1 may be affected by Christmas/familiarity |
|  | Total |  | 29 |  |


| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 5(a)(i) | $\mathrm{z}_{1}=(236.5-244.43) / 4.09$ | B1 |  | B1 attempt to find tail probability < |
|  | $=-1.939$ | B1 |  | B1 Use of 236.5 as upper bound of class or equivalent |
|  | $\begin{aligned} \text { probability }<236.5 & =1-0.9737 \\ & =0.0263 \end{aligned}$ | M1 |  | M1 method for probability not dependent on B marks - |
|  | expected number in first class $=$ $0.0263 \times 105=2.76$ expected number in last class | m1 |  | m1 their prob $\times 105$ |
|  | $\begin{aligned} & 105-2.76-9.21-21.48-29.93- \\ & 24.89-12.36=4.37 \end{aligned}$ | M1 | 5 | M1 method for E last class |
|  | $\begin{gathered} {[\text { or } \mathrm{z}=(251.5-244.43) / 4.09} \\ =1.729 \\ \text { probability }>251.5=1-0.9581 \\ =0.0419 \\ \text { expected number in last class }= \\ 0.0419 \times 105=4.40] \end{gathered}$ |  |  |  |
| (ii) | Combining classes where $\mathrm{E}<5$ O E |  |  |  |
|  | $\begin{array}{lll}<239 & 12 & 11.97\end{array}$ | M1 |  | M1 attempt to combine classes |
|  | 240-242 18 21.48 | m1 |  | m 1 correct method for combining |
|  | 243-245 37 29.93 |  |  | classes + correct classes combined |
|  | $\begin{array}{ccc} 246-248 & 21 & 24.89 \\ >248 & 17 & 16.73 \end{array}$ |  |  |  |
|  | $\mathrm{H}_{0}$ : Normal distribution is adequate model <br> $\mathrm{H}_{1}$ : Normal distribution is not adequate model |  |  |  |
|  | $\begin{aligned} & \Sigma(\mathrm{O}-\mathrm{E})^{2} / \mathrm{E}=0.03^{2} / 11.97+ \\ & 3.48^{2} / 21.48+7.07^{2} / 29.93+ \end{aligned}$ | M1 |  | M1 attempt at $\Sigma(\mathrm{O}-\mathrm{E})^{2} / \mathrm{E}-$ their Es |
|  | $3.89^{2} / 24.89+0.27^{2} / 16.73=2.85$ | A1 |  | A1 2.85 (2.8~2.9) needs previous M1m1M1 |
|  | c.v. $\chi_{2}{ }^{2}$ is 4.605 | B1ヶ |  | B1 $\checkmark 2 \mathrm{df}$ |
|  |  | B1 |  | B1 4.605 |
|  | No significant evidence that the normal distribution is not an adequate model for the temperature at which the lubricant becomes ineffective. | $\mathrm{A} 1 \checkmark$ <br> A1 $\checkmark$ | 8 | A $1 \checkmark$ conclusion - needs all previous method marks; 5 marks for first (a)(i) and comparison with upper tail of $\chi^{2}$ A1 $\sqrt{ }$ in context |
| (b) | Kabeera's claim is correct as this is a large sample. Mean will be | E1 |  | E1 claim correct |
|  | approximately normally distributed whether the underlying distribution is normal or not. The sample will also give a good estimate of the standard deviation | E1 | 2 | E1 large sample/ central limit theorem |
|  | Total |  | 15 |  |
|  | TOTAL |  | 75 |  |


[^0]:    Further copies of this Mark Scheme are available from: aqa.org.uk

