

General Certificate of Education (A-level)
June 2011

Statistics SS03

(Specification 6380)

Statistics 3

Final

Mark Scheme

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Key to mark scheme abbreviations

| M | mark is for method |
|-------------|--|
| m or dM | mark is dependent on one or more M marks and is for method |
| A | mark is dependent on M or m marks and is for accuracy |
| В | mark is independent of M or m marks and is for method and accuracy |
| E | mark is for explanation |
| √or ft or F | follow through from previous incorrect result |
| CAO | correct answer only |
| CSO | correct solution only |
| AWFW | anything which falls within |
| AWRT | anything which rounds to |
| ACF | any correct form |
| AG | answer given |
| SC | special case |
| OE | or equivalent |
| A2,1 | 2 or 1 (or 0) accuracy marks |
| –x EE | deduct x marks for each error |
| NMS | no method shown |
| PI | possibly implied |
| SCA | substantially correct approach |
| c | candidate |
| sf | significant figure(s) |
| dp | decimal place(s) |

No Method Shown

Where the question specifically requires a particular method to be used, we must usually see evidence of use of this method for any marks to be awarded.

Where the answer can be reasonably obtained without showing working and it is very unlikely that the correct answer can be obtained by using an incorrect method, we must award **full marks**. However, the obvious penalty to candidates showing no working is that incorrect answers, however close, earn **no marks**.

Where a question asks the candidate to state or write down a result, no method need be shown for full marks.

Where the permitted calculator has functions which reasonably allow the solution of the question directly, the correct answer without working earns **full marks**, unless it is given to less than the degree of accuracy accepted in the mark scheme, when it gains **no marks**.

Otherwise we require evidence of a correct method for any marks to be awarded.

| SS03 O | Solution | Marks | Total | Comments |
|-----------|--|-------|-------|---|
| 1(a) | H ₀ Samples are taken from identical populations H ₁ Samples are not taken from identical populations (or population average substance level differs) | B1 | 1000 | Or equivalent hypotheses referring to population medians. Allow pop median A/B or η_A and η_B Must have 'average'; disallow mean |
| | 2 tail 5% | | | |
| | Ranks A 1 4 5 6 8 10.5 14 15 (15) (12) (11) (10) (8) (5.5) (2) (1) | M1 | | Attempt at ranks as one group - either way |
| | B 2 3 7 9 10.5 12 13 (14) (13) (9) (7) (5.5) (4) (3) | | | For ties |
| | $T_A = 1 + 4 + \dots + 15 = 63.5$ (or 64.5) $T_B = 2 + 3 + \dots + 13 = 56.5$ (or 55.5) | M1 | | Totals (dep on ranks - any) |
| | $U_A = 63.5 - \frac{8 \times 9}{2} = 27.5$ | | | |
| | $U_B = 56.5 - \frac{7 \times 8}{2} = 28.5$ Test stat $U = 27.5$ | m1A1 | | U calculated - either correct |
| | -27.5 | | | |
| | cv = 13 | B1 | | For cv |
| | U >13 | A1 | | Correct U chosen for comparison with relevant cv (27.5/13 or 28.5/43) |
| | Accept H _o | A1 | | Only if ts/cv correct |
| | No significant evidence at the 5% level to suggest that there is any difference in the average level of the substance for drugs A and B. | E1 | 10 | In context (can ft) |
| (b) | Type II error is to accept H_0 when actually H_0 is not true. | B1 | | Allow 'my conclusion was wrong' |
| | This would mean that the conclusion to the test in part (a) that there is no significant difference in the average level of the substance for drugs A and B is incorrect and there is actually a difference between the two. | E1 | 2 | For context 2 marks only if in context and correct |
| | Total | | 12 | |

| Q | Solution | Marks | Total | Comments |
|---------|--|-------|-------|--|
| 2(a)(i) | $H_0 \ \mu, \ \eta = 10.8$ | | | |
| () () | • | B1 | | Or equivalent in words, eg 'average time |
| | $H_1 \mu, \eta > 10.8$ | Бī | | spent on study' or 'population average' |
| | 1 tail 5% | | | spent on study of population average |
| | diffs 6.65 3.85 1.5 0.8 -3.7 | | | |
| | rank 10 6 2 1 5 | M1 | | For differences from 10.8 |
| | Tunk 10 0 2 1 3 | | | |
| | diffs 4.35 5.4 -3.2 -4.05 -2.2 | m1 | | Ranks as one group dep on differences |
| | rank 8 9 4 7 3 | | | (allow either way) SC1 for sign test |
| | | | | |
| | $T_{+} = 10 + 6 + 2 + 1 + 8 + 9 = 36$ | m1 | | Total of any ranks dep on diffs |
| | $T_{-} = 5 + 4 + 7 + 3 = 19$ | A1 | | One correct |
| | Test stat $T = 19$ | | | |
| | n = 10 | | | |
| | cv = 11 | B1 | | For cv |
| | T > 11 | m1 | | Comparison lower (plausible) T (not –ve) |
| | | | | and cv. Can ft or 44/36 |
| | Accept H ₀ | A1 | | |
| | | | | |
| | There is no significant evidence to suggest | Г1 | 0 | T |
| | that average time spent per week of term has increased from 10.8 hours. | E1 | 9 | In context |
| | nas mercasea from 10.5 nours. | | | |
| (a)(ii) | Conclusions cannot be generalised to | | | |
| | whole population. | E1 | 1 | For any one point clearly explained |
| | Students at the college concerned may not | | | (not 'may have lied', 'not correctly |
| | represent a random sample of all such | | | recorded') |
| | students in the country. | | | |
| | Study patterns may vary at different times of the year. | | | |
| | of the year. | | | |
| (b)(i) | Wilcoxon signed-rank takes takes into | | | |
| | account the magnitude of the ranks of the | | | |
| | differences whereas the sign test only | | | |
| | considers the sign of those differences. | E1 | 1 | Or 'magnitude of differences' |
| | or Wilcoxon signed-rank is more likely | | | (not 'takes data/size of data into account') |
| | to detect a difference if one exists. | | | |
| | or More powerful. | | | |
| (ii) | If a direction/preference only was given | | | |
| | then there would be no numerical data | | | |
| | available to find the differences in the | В1 | | For one valid situation - |
| | data that need to be used for the Wilcoxon | | | a direction/preference or asymmetrical |
| | signed-rank test. An example would be if | | | |
| | students only had to state whether they | | | |
| | were studying more hours, less hours or | E1 | 2 | explained clearly in context |
| | the same hours this year as last year. | | | |
| | or If data to be analysed was very asymmetrical. An example could be that | | | |
| | the times for study were found to be skew. | | | |
| | Total | | 13 | |
| | Total | | 15 | |

| 2()() | | Solution | | | Marks | Total | | Con | nments | |
|---------|--|--|---|----------|----------------|--------------------------------|--|------------------------------|---------------------------|-------------------------|
| 3(a)(i) | 1 2-4 > 4 | None 5.74 18.64 13.62 | 1 or 2 > 2 7.40 2.8 24.04 9.3 17.57 6.8 | 7 2 | M1 A1 A1 | 3 | Method for (16 × 38 / 4 correct; All correct | 106 etc); allow slig | ft incorre | ct totals |
| (ii) | There is an below 5. | n expected | d frequency that | is | E1 | 1 | Must refe | r to expec | ted freque | ency |
| (iii) | "More than 2 falls" pooled with "1 or 2 falls" to make 2 categories of fall: "None" or "One or more" | | | E1 | 1 | Allow '1' categories and "More | of medic | | to make 2 en: "1 to 4" | |
| (iv) | number of | f medicati per of) fall | s is not indepen | | B1 | | | Associat | ion | |
| | Obs 1 2-4 More 4 | Nor 10 21 7 | 6 | | M1 | | Obs 1 - 4 More 4 | None 31 7 | 1 or 2 31 18 | More 2 6 13 M1 |
| | Exp 1 2–4 More 4 | Nor 5.7-18.6 | 4 10.26 54 33.56 | <u> </u> | m1 | | For 3E co | rrect None | 1 or 2 | More 2 |
| | $ts = \sum \frac{(C)}{1}$ $\frac{4.26^2}{5.74} + \frac{4}{1}$ | $\frac{1.26^2}{0.26} + \frac{2.3}{18}$ | | | m1 | | 1-4 More 4 For ts; Ya | 24.38 13.62 tes used N | 31.43 17.57 40 | 12.19 6.81 m1 |
| | $\frac{2.36^{2}}{33.36} + \frac{6}{1}$ $ts = 10.4$ (SC4 ts = df = 2 1 | (10.0–11 10.4 NM | .0) S if (i) ok) | | A1 | | ts = 13.80 (SC4 ts = | | | |
| | ts > 9.21 | 0 | - | | B1 | | | | | |
| | | nce to sug | gest that the num lent of number o | | E1 | 7 | E1 only if sense Allow B1 10.597 | | | |

SS03 (cont)

| SS03 (cont) | 914 | N/ 1 | /T 4 1 | C . |
|-------------|---|----------------|--------|---|
| Q | Solution | Marks | Total | Comments |
| 3(b)(i) | No falls At least one 0 1680 660 1 448 180 2 392 192 3+ 280 168 | M1 m1 A1 | 3 | Method for frequencies eg 0.60×2800 4 correct all correct |
| (ii) | H_0 (Number of) falls is independent of number of chronic diseases suffered H_1 (Number of) falls is not independent of number of chronic diseases suffered $\sum (O-E)^2$ | B1 | | OE: H ₀ No association H ₁ Association |
| | $ts = \sum \frac{(O-E)^2}{E} = 18.4$ $df = 3$ 1% $ev = 11.345$ $ts > 11.345$ | B1 M1 | | For 11.345 Comparison with 18.4 Allow M1 for 6.251, 7.815, 9.348, 12.838 |
| | Reject H ₀ Significant evidence to suggest that the number of falls suffered is not independent of number of diseases suffered | E1 | 4 | Conclusion in context |
| (iii) | Must have attempted (i) and (ii) to gain marks in (iii) | | | E0,0 if conclusion in (ii) is Accept H ₀ |
| | Women who do not suffer from any chronic diseases are less likely than expected to have a fall. | E1 | | |
| | Women who suffer from 3 or more chronic diseases are more likely than expected to have a fall. | E1 | 2 | Or equivalent – no need to refer to expected frequencies but disallow comment referring simply to a comparion |
| | [Women who suffer from 2 chronic diseases are more likely than expected to have a fall.] | | 21 | of observed frequencies |
| | Total | | 21 | |

SS03 (cont)

| Q Q | Solution | Marks | Total | Comments |
|--------------|--|---------|-------|---|
| 4(a) | r = 0.895 | B3 | 3 | SC2 0.89/0.90/0.894 SC1 0.9 |
| 4(a) | 7 – 0.073 | DS | 3 | Allow M1 summations |
| | | | | M1 correct use of S_{xx} S_{xy} S_{yy} |
| | | | | THE CONTEST USE OF S _{XX} S _X y S _{yy} |
| (b) | $H_0 \rho = 0$ | B1 | | OE in words |
| | $H_1 \rho > 0$ | B1 | | H_0 pop PMCC = 0 or |
| | 17 | | | H ₀ no association between BMR and BMI |
| | | | | H ₁ correct direction B1 |
| | | | | |
| | ts $r = 0.895$ | | | |
| | n = 10 cv = 0.5494 | B1 | | CAO for cv |
| | r > 0.5494 reject H ₀ | M1 | | ft provided $-1 < r < +1$ |
| | | | | |
| | Significant evidence that there is a | A1 | 5 | For Reject H ₀ ts/cv correct |
| | positive correlation between BMR and | | | |
| | BMI. | | | |
| | Men with a higher BMR tend to have a | | | |
| | higher BMI. | | | |
| (c) | Ranks for BMR | | | |
| | 10 9 8 7 6 5 4 21/2 21/2 1 | | | |
| | or 1 2 3 4 5 6 7 8½ 8½ 10 | M1 | | Ranks (can be reversed) |
| | | m1 | | Ties |
| | | A1 | | For d 2 1 2 0 3 0 3 ½ ½ 3 |
| | | | | $\sum d^2 = 4 + 1 + \dots + 9 = 36\frac{1}{2}$ |
| | | | | |
| | CDCC 0.770/11000\ : : | D2 | ~ | $6 \times 36 \frac{1}{2} - 0.770$ |
| | SRCC $r_s = 0.778(11909)$ ignore sign | B2 | 5 | SRCC = $1 - \frac{6 \times 36 \frac{1}{2}}{10 \times 99} = 0.779$ |
| | NMS SC4 $r = 0.78$ SC2 $r = 0.8$ | | | Reversed ranks $\sum d^2 = 292.5 \ r = -0.773$ |
| | SC0 $r = 0.5636$ | | | M1A1 must be 0.779/–0.773 |
| | SC3 $r = 0.770/0.769$ with ranks | | | WITAT must be 0.7797-0.773 |
| | SC2 $r = 0.770/0.769$ no ranks | | | |
| | | | | |
| (d) | There is a significant positive correlation | | | |
| | between BMR and BMI and there is | | | |
| | strong positive rank correlation between | E1 | | Both results put together |
| | BMR and level of daily physical activity. | (no ft) | | |
| | Managha hara - 12-1 DMT - 1-1 | | | |
| | Men who have a high BMI tend to have a | E1 | 2 | Interpretation in context (not just repeat of |
| | high BMR as do men who have a high level of daily physical activity. | EI | 2 | conclusion) |
| | iever or dairy physical activity. | | | Conclusion) |
| (e)(i) | BMR and BMI measurements are | | | Mention of normal distribution or linear |
| | normally (or bivariate normal) distributed | B1 | 1 | relationship seen |
| | , | | | _ |
| (ii) | Ranks only available for level of daily | | | |
| | physical activity so SRCC is the only | | | |
| | correlation coefficient that can be evaluate | E1 | 1 | Clearly in context |
| | or No actual values given for DPA | | | |
| | Total | | 17 | |

SS03 (cont)

| Q | Solution | Marks | Total | Comments |
|------|--|----------|-------|---|
| 5(a) | The purpose is to ensure that each participant has the same opportunity to be | | | 'Student effect' eliminated |
| | assigned to any one of the three methods so that each method group should be | E1 | | 'bias eliminated' |
| | roughly equivalent. Therefore any difference observed between method | | | |
| | groups can be linked to the effect of the method, not due to a characteristic of the | E1 | 2 | More likely to detect any difference as groups more equivalent |
| | individuals in the group. | | | E0 'more accurate' |
| (b) | H₀ Samples from identical populations H₁ Samples not from identical populations | B1 | | Or hypotheses referring to difference between at least 2 population averages (not mean) |
| | Ranks Prog Comp Cont 5 13 17 1 15 3 10 8 16 2 11 7 | M1 | | For ranks as one group (can be reversed) |
| | 6 12 14 4 8 10 1 17 12 6 4 14 13 5 7 11 3 15 9 9 2 16 | | | |
| | Totals of ranks: | | | |
| | $T_{\text{prog}} = 44/64$ $T_{\text{comp}} = 66/24$ $T_{\text{cont}} = 43/65$ $n_{\text{prog}} = 6$ $n_{\text{comp}} = 5$ $n_{\text{cont}} = 6$ | m1 A1 | | Totals ft but dep on ranks 2 totals correct |
| | $\sum_{i=1}^{m} \frac{T_i^2}{n_i} = \frac{44^2}{6} + \frac{66^2}{5} + \frac{43^2}{6} = 1502.03$ or $322.7 + \dots \qquad \text{m1 implied}$ or $682.7 + \dots$ | m1 m1 | | Numerators correct Denominators correct |
| | $H = \frac{12}{17 \times 18} \times 1502.03 - (3 \times 18)$ | m1 | | H formula correctly used (need a $\sum_{i=1}^{m} \frac{T_i^2}{n_i}$) |
| | = 4.90 | A1 | | 4.80–5.00 |
| | Critical value from $\chi_2^2 = 5.99$ | В1 | | For cv (correct cv only) |
| | H < 5.99 | | | |
| | Accept H ₀ No reason to doubt that samples are from identical poulations. No significant difference in average scores in test for the 3 methods. | E1 | 10 | Conclusion correct in context (must have ts/cv both correct) |
| | Total | | 12 | |
| | TOTAL | | 75 | |